

# Stable Limit Theorem for $U$ -Statistics Processes Indexed by a Random Walk

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# Outline

## Introduction

Random Walks in Random Scenery  
 $U$ -Statistics

## Main Results

Limit Theorems  
Sketch of Proofs

# Partial Sum Processes

for  $(\xi_n)_{n \in \mathbb{N}}$  iid with  $E\xi_i = 0$  and  $E\xi_i^2 < \infty$

$$\left( \frac{1}{\sqrt{n}} \sum_{i=1}^{\lfloor nt \rfloor} \xi_i \right)_t \Rightarrow B$$

for a Brownian motion  $B$  (self-similar process with index  $1/2$ )

different scaling and different self-similarity of the limes process for

- ▶ heavy tailed random variables
- ▶ long range dependent random variables

# Definition

**random walk:**  $(X_n)_{n \in \mathbb{N}}$  iid  $\mathbb{Z}$ -valued random variables in the normal domain of attraction of an  $\alpha$ -stable law  $F_\alpha$  with  $0 < \alpha \leq 2$

$$S_n := \sum_{m=1}^n X_m$$

**scenery:**  $(\xi(i))_{i \in \mathbb{Z}}$  iid  $\mathbb{R}$ -valued random variables in the normal domain of attraction of an  $\beta$ -stable law  $F_\beta$  with  $0 < \beta \leq 2$

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**random walk in random scenery:** partial sum process

$$W_n(t) = \sum_{i=1}^{\lfloor nt \rfloor} \xi(S_i)$$

(or linear interpolated version)

# Limit Theorem:

## Theorem (Kesten, Spitzer, 1979)

For  $1 < \alpha \leq 2$  in  $C[0, 1]$

$$n^{-1+\frac{1}{\alpha}-\frac{1}{\alpha\beta}} W_n(t) \Rightarrow \Delta_t$$

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### properties of $\Delta$

- ▶ self-similar with index  $\delta = 1 - \frac{1}{\alpha} + \frac{1}{\alpha\beta}$
- ▶ non Gaussian even if  $\beta = 2$  (scenery has Gaussian limit)
- ▶ continuous even if  $\beta < 2$  (scenery has limit with jumps)

# Definition of Limit Process:

Rewrite  $W_n = W_n(1)$  as

$$W_n = \sum_{i=1}^n \xi(S_i) = \sum_{x \in \mathbb{Z}} N_n(x) \xi(x)$$

with occupation times

$$N_n(x) = \sum_{i=1}^n \mathbb{1}_{\{S_i=x\}}$$



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$$N_n(x) = \sum_{i=1}^n \mathbb{1}_{\{S_i=x\}}$$

$$\Delta_t = \int_{-\infty}^{\infty} L_t(x) dZ(x)$$

with two-sided  $\beta$ -stable process  $Z$ , local time  $L$  of an  $\alpha$ -stable process  $S^*$ , i.e.

$$\int_0^t \mathbb{1}_{[a,b)}(S_s^*) ds = \int_a^b L_t(x) dx$$

# Transient Random Walks

## Theorem (Castell, Guillin-Plantard, Pène, 2013)

If  $\alpha = 1$ ,  $\beta \neq 1$ , then in  $D[0, 1]$  (equipped with  $M_1$  topology)

$$\left( n^{-\frac{1}{\beta}} (\log n)^{\frac{1-\beta}{\beta}} W_n(t) \right)_t \Rightarrow C_1 Z$$

- ▶  $\beta$ -stable process  $Z$  (discontinuous)
- ▶  $S_i = S_j$  less often, so less dependence

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## Theorem (Castell, Guillin-Plantard, Pène, 2013)

If  $\alpha < 1$ , then

$$\left( n^{-\frac{1}{\beta}} W_n(t) \right)_t \Rightarrow C_2 Z$$

# Definition

- ▶  $(X_n)_{n \in \mathbb{N}}$  stationary sequence of random variables
- ▶  $h : \mathbb{R}^2 \rightarrow \mathbb{R}$  measurable and symmetric

## Definition

The (bivariate) *U-statistic*  $U_n(h)$  with kernel  $h$  is defined as

$$U_n(h) := \sum_{1 \leq i < j \leq n} h(X_i, X_j).$$

# Examples

1.  $h(x, y) = |x - y|$

$$\frac{2}{n(n-1)} U_n(h) = \frac{2}{n(n-1)} \sum_{1 \leq i < j \leq n} |X_i - X_j|,$$

Gini's mean difference

2.  $h(x, y) = \frac{1}{2}(x - y)^2$

$$\frac{2}{n(n-1)} U_n(h) = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2.$$

sample variance

# Hoeffding-Decomposition

$U_n(h)$  can be decomposed into a *linear part* and a *degenerate part*

$$U_n(h) = \frac{n(n-1)}{2}\theta + (n-1) \sum_{i=1}^n h_1(X_i) + U_n(h_2)$$

with

$$\theta := Eh(X_1, X_2)$$

$$h_1(x) := Eh(x, X_2) - \theta$$

$$h_2(x, y) := h(x, y) - h_1(x) - h_1(y) - \theta.$$

Note that  $h_2$  is degenerate, i.e. for  $i < j$  and  $X_i, X_j$  independent

$$E[h_2(X_i, X_j)|X_i] = 0.$$

# Hoeffding-Decomposition II

Hence for independent data and  $i < j < k$

$$\begin{aligned} E[h_2(X_i, X_j)h_2(X_i, X_k)] &= E[E[h_2(X_i, X_j)h_2(X_i, X_k)|X_i, X_j]] \\ &= E[h_2(X_i, X_j)E[h_2(X_i, X_k)|X_i, X_j]] = 0 = E[h_2(X_i, X_j)] \cdot E[h_2(X_i, X_k)] \end{aligned}$$

So the summands of  $U_n(h_2)$  are uncorrelated and we have

$$\text{Var } U_n(h_2) = O(n^2)$$

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## Theorem (Hoeffding, 1948)

*If  $(X_n)_{n \in \mathbb{N}}$  is a sequence of iid random variables and  $\text{Var } h_1(X_1) < \infty$ , then*

$$n^{-\frac{3}{2}} (U_n(h) - \theta) \xrightarrow{\mathcal{D}} N(0, \text{Var } h_1(X_1)).$$



# Existing Results

## under second moments (or higher) and dependence:

- ▶ Sen (1963):  $m$ -dependence
- ▶ Yoshihara (1976): absolute regularity
- ▶ Dehling, Taqqu (1989): long-range dependent Gaussian processes
- ▶ Guillin-Plantard, Ladret (2005): random walk in random scenery
- ▶ Dehling, Wendler (2010): strong mixing

## without second moments, under independence:

- ▶ Heinrich, Wolf (1993)
- ▶ Dabrowski, Dehling, Mikosch, Sharipov (2002)

# $U$ -Statistic indexed by a random walk

$$U_n = \sum_{1 \leq i < j \leq n} h(\xi(S_i), \xi(S_j))$$

Assumption:  $E|h(\xi(1), \xi(2))| < \infty$ ,  $E|h(\xi(1), \xi(1))| < \infty$   
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## Hoeffding decomposition

$$U_n = (n-1) \sum_{i=1}^n h_1(\xi(S_i)) + \sum_{1 \leq i < j \leq n} h_2(\xi(S_i), \xi(S_j))$$

with  $h_1(x) := E(h(x, \xi(1)))$

$h_2(x, y) := h(x, y) - h_1(x) - h_1(y)$ .

# Weak Convergence

- ▶  $h_1(\xi(1))$  in normal domain of attraction of  $\beta$ -stable law,  $1 < \beta \leq 2$
- ▶  $E|h(\xi(1), \xi(2))|^\eta < \infty$  with  $\eta = \frac{2\beta'}{1+\beta'}$  for a  $\beta' > \beta$
- ▶  $X_1$  in normal domain of attraction of an  $\alpha$ -stable law  $F_\alpha$

## Theorem (Franke, Pène, W.)

If  $1 < \alpha \leq 2$ , then in  $C[0, 1]$

$$\left( n^{-2+\frac{1}{\alpha}-\frac{1}{\alpha\beta}} U_n(t) \right)_{t \in [0,1]} \Rightarrow (\Delta t)_{t \in [0,1]},$$

$U_n(t) = U_k$  if  $nt = k$ , linear interpolated in between

# Transient Random Walks

## Theorem (Franke, Pène, W.)

If  $\alpha = 1$ , then in  $D[0, 1]$  (equipped with  $M_1$  topology)

$$\left( n^{1-\frac{1}{\beta}} (\log n)^{\frac{1-\beta}{\beta}} U_{\lfloor nt \rfloor} \right)_{t \in [0, 1]} \Rightarrow (C_3 Z_t)_{t \in [0, 1]},$$

$Z$   $\beta$ -stable process which is the limit of  $n^{-1/\beta} \sum_{i=1}^{\lfloor nt \rfloor} h_1(\xi_i)$

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## Theorem (Franke, Pène, W.)

If  $0 < \alpha < 1$ , then in  $D[0, 1]$

$$\left( n^{-1-\frac{1}{\beta}} U_{\lfloor nt \rfloor} \right)_{t \in [0,1]} \Rightarrow (C_4 Z_t)_{t \in [0,1]},$$

# Law of the iterated logarithm

- ▶  $E|h_1(\xi(i))|^p < \infty$
- ▶  $E|X_i|^p < \infty$  for some  $p > 2$

Theorem (Franke, Pène, W.)

$$\limsup_{n \rightarrow \infty} \frac{U_n}{n^{\frac{7}{4}} (\log \log n)^{\frac{3}{4}}} = \frac{2^{\frac{1}{4}} \text{Var}(h_1(\xi(1)))^{\frac{1}{2}}}{3 \text{Var}(X)^{\frac{1}{4}}}$$
$$\liminf_{n \rightarrow \infty} \frac{U_n}{n^{\frac{7}{4}} (\log \log n)^{\frac{3}{4}}} = -\frac{2^{\frac{1}{4}} \text{Var}(h_1(\xi(1)))^{\frac{1}{2}}}{3 \text{Var}(X)^{\frac{1}{4}}}$$

*almost surely*

# Lemma for Degenerate Part

## Hoeffding decomposition

$$U_n = (n-1) \sum_{i=1}^n h_1(\xi(S_i)) + \sum_{1 \leq i < j \leq n} h_2(\xi(S_i), \xi(S_j))$$

## Lemma

for  $0 < \alpha \leq 2$

$$\max_{k \leq n} \sum_{1 \leq i < j \leq k} h_2(\xi(S_i), \xi(S_j)) = o(n^{2 - \frac{1}{\alpha'} + \frac{1}{\alpha'\beta}})$$

almost surely with  $\alpha' := \max(1, \alpha)$



# Truncation of Kernel

$$h_{0,n}(x, y) := \begin{cases} h(x, y) & \text{if } |h(x, y)| \leq a_n = n^{\frac{1+\beta'}{\alpha'\beta'}} \\ 0 & \text{else} \end{cases}.$$

Hoeffding decomposition of truncated kernel

$$\begin{aligned} h_{1,n}(x) &:= E(h_{0,n}(x, \xi(1))) \\ h_{2,n}(x, y) &:= h_{0,n}(x, y) - h_{1,n}(x) - h_{1,n}(y). \end{aligned}$$

$$\sum_{1 \leq i < j \leq n} h_{0,n}(\xi(S_i), \xi(S_j)) = (n-1) \sum_{i=1}^n h_{1,n}(\xi(S_i)) + \sum_{1 \leq i < j \leq n} h_{2,n}(\xi(S_i), \xi(S_j))$$

# Hoeffding Decomposition Read Backwards

$$\begin{aligned}
 \sum_{1 \leq i < j \leq k} h_2(\xi(S_i), \xi(S_j)) &= \sum_{1 \leq i < j \leq k} h(\xi(S_i), \xi(S_j)) - (n-1) \sum_{i=1}^n h_1(\xi(S_i)) \\
 &= \left( \sum_{1 \leq i < j \leq k} h(\xi(S_i), \xi(S_j)) - \sum_{1 \leq i < j \leq k} h_{0,n}(\xi(S_i), \xi(S_j)) \right) \\
 &\quad - (n-1) \sum_{i=1}^n h_1(\xi(S_i)) + \sum_{1 \leq i < j \leq k} h_{0,n}(\xi(S_i), \xi(S_j)) \\
 &= \sum_{1 \leq i < j \leq k} (h(\xi(S_i), \xi(S_j)) - h_{0,n}(\xi(S_i), \xi(S_j))) \\
 &\quad - (n-1) \sum_{i=1}^n (h_1(\xi(S_i)) - h_{1,n}(\xi(S_i))) + \sum_{1 \leq i < j \leq k} h_{2,n}(\xi(S_i), \xi(S_j))
 \end{aligned}$$

# Truncation of Linear Part

$$E|h_1(\xi(1)) - h_{1,n}(\xi(1))| \leq \frac{1}{a_n^{\eta-1}} E|h(\xi(1), \xi(2))|^\eta.$$

with triangular inequality

$$E \left| (n-1) \sum_{i=1}^n (h_1(\xi(i)) - h_{1,n}(\xi(i))) \right| \leq Cn^2 \frac{1}{a_n^{\eta-1}} = o(n^{2-\frac{1}{\alpha'} + \frac{1}{\alpha'\beta}})$$

with Markov inequality

$$(n-1) \sum_{i=1}^n (h_1(\xi(i)) - h_{1,n}(\xi(i))) = o_p(n^{2-\frac{1}{\alpha'} + \frac{1}{\alpha'\beta}})$$

similar: truncation of  $U_n$

# Degenerate Part

$$E(h_{2,n}(\xi(1), \xi(2)))^2 \leq C\alpha_n^{2-\eta} E|h_2(\xi(1), \xi(2))|^\eta = Cn^{\frac{2}{\alpha'\beta'}}$$

as summands are uncorrelated

$$E\left(\sum_{\substack{1 \leq i < j \leq k \\ S(i) \neq S(j)}} h_{2,n}(\xi(S_i), \xi(S_j))\right)^2 \leq Cn^{\frac{2}{\alpha'\beta'}} E\left(\sum_{x \in \mathbb{Z}, y \in \mathbb{Z}} N_n^2(x) N_n^2(y)\right) \\ = o(n^{4 - \frac{2}{\alpha'} + \frac{2}{\alpha'\beta'}})$$

because for occupation times  $N_n(x)$

$$E\left(\sum_{x \in \mathbb{Z}} N_n^2(x)\right)^2 = O(n^{4 - \frac{2}{\alpha}} \log^2 n)$$

# Linear Part

$$(n-1) \sum_{i=1}^n h_1(\xi(S_i))$$

weak convergence

- ▶  $1 < \alpha \leq 2$ : Kesten, Spitzer (1979)
- ▶  $0 < \alpha \leq 1$ : Castell, Guillin-Plantard, Pène (2013)

law of the iterated logarithm

- ▶  $\alpha = \beta = 2$ : Zhang (2001)

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## Open Questions:

- ▶ limit for  $\beta \leq 1$  (Hoeffding decomposition not possible)
- ▶ limit if  $h_1 = 0$  (degenerate  $U$ -statistic)

# Acknowledgement

Thank you for listening!

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